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Process optimization of *Syringa oblata* Lindl. by response surface methodology and its effect on *Staphylococcus xylosus* biofilm†

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Syringa oblata Lindl. (*S. oblata*) is a medicinal plant with effective broad-spectrum antibacterial activity, which can also inhibit *Streptococcus suis* biofilm formation. The processing of herbal medicine can purify medicinal materials, provide acceptable taste, reduce toxicity, enhance efficacy, influence performance and facilitate preparation. Thus, the aim of this study was to enhance the biofilm inhibition activity of *S. oblata* toward *Staphylococcus xylosus* (*S. xylosus*) using the best processing method. The content of rutin and flavonoids and the ability to inhibit the biofilm formation by *S. oblata* were examined using four processing methods. One of the best methods, the process of stir-frying *S. oblata* with vinegar, was optimized based on the best rutin content by response surface methodology. The histidine content and *hisB* gene expression of *S. xylosus* biofilm *in vitro*, resulting from stir-frying *S. oblata* with vinegar, were evaluated and were found to be significantly decreased and down-regulated, respectively. The results show that *S. oblata* stir-fried with vinegar can be used to effectively treat diseases resulting from *S. xylosus* infection. This is because it significantly inhibited *S. xylosus* biofilm formation by interfering with the biosynthesis of histidine; thus, its mechanism of action is decreasing histidine synthesis.

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1. Introduction

The processing of herbal medicines by different methods, including parching, stir-baking with liquid, stir-frying with liquid, calcination, roasting in fresh cinders, steaming and boiling,¹ plays an important role² in achieving unique functions such as toxicity reduction, improving the pharmacological properties,³ stabilizing active ingredients,⁴ and flavor improvement or correction and content purification before clinical application.³ The safety, effectiveness, and quality of processed herbal medicines have attracted worldwide attention.² Generally, the main mechanisms underlying herbal processing have been found to be related to the changes in the composition and/or activity of the components in the herbs.³

Syringa oblata Lindl. (*S. oblata*) belongs to the Oleaceae plant family. The leaves and bark have a bitter taste and have

important medicinal and industrial application purposes.⁵ *S. oblata* has promising biological effects such as anti-tumor, anti-hypertensive, anti-oxidant, anti-inflammatory activities.⁵ One of the major constituents associated with the anti-bacterial activity of *S. oblata* is rutin,⁶ which has been reported as the main bioactive ingredient in *S. oblata* that inhibits the formation of biofilm in *Streptococcus suis* (*S. suis*).⁷ However, it is unknown whether the concentration of rutin in *S. oblata* is affected by different processing methods and the effect of *S. oblata* against the formation of biofilm by *Staphylococcus xylosus* (*S. xylosus*) remains unexplored. The application of traditional Chinese medicine processing has a long history, and modern research shows that it has a certain scientific foundation. In order to improve the applications of *S. oblata*, it is important to influence its composition and pharmacological effects.

S. xylosus, a member of the coagulase-negative staphylococci (CoNS)⁸ is the most frequently isolated species from cows with mastitis⁹ and has the ability to form a biofilm.^{10,11} L-histidine biosynthesis is an ancient metabolic pathway present in bacteria and plants¹² and is involved in the formation of biofilm in *S. xylosus*¹³ and *Saccharomyces cerevisiae*.¹⁴ However, there are no relevant pharmacological studies that evaluate the antibacterial activity of raw and processed *S. oblata*. Thus, the selection and optimization of processing methods may be able to

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effectively improve the association between the main components in herbal medicines and anti-bacterial activity.

The conventional optimization of a processing method in analytical chemistry involves monitoring the influence of one factor at a time on an experimental response, while other parameters are kept at a constant level. This, however, has certain disadvantages in that it does not depict all the effects of the parameter on the response, and the increase in the number of experiments necessary to conduct the research leads to an increase in the amount of time and expenses as well as an increase in the consumption of reagents and materials.^{15,16} Thus, this process response is a direct function of the single varied parameter.¹⁶ In order to overcome this problem, statistical optimization procedures allow one to take the interaction of variables into consideration by using multivariate statistical techniques.¹⁶ Currently, response surface methodology (RSM) is one of the most conventional and potent approaches to multivariate optimization analytical methodologies.^{17,18} This statistical tool tests the applicability and reliability of optimization to improve the process,^{19,20} reduce the number of experiments required to result in a less laborious task,²¹ and it indicates the possible influences of some variables on others.¹⁸ RSM is faster and more informative than the classical one-variable-at-a-time approach or the use of full factorial designs.²² Furthermore, RSM has been successfully used for developing, improving and optimizing processes.¹⁶

In this study, the changes in the rutin and flavonoid content and the effect of *S. oblata* against *S. xylosus* biofilm formation *in vitro* were tested under different *S. oblata* processing methods. The best method of processing *S. oblata* was then optimized using RSM. The histidine content and *hisB* gene expression of *S. xylosus* biofilm *in vitro* were evaluated after stir-frying *S. oblata* with vinegar. This study provides an explanation of how the methods of processing *S. oblata* affect *S. xylosus* biofilm formation and lays the foundation for further developing and optimizing the use of processed *S. oblata*. The detailed screening workflow is shown in Fig. 1.

2. Materials and methods

2.1. Plant materials

S. oblata, identified by Professor Xiuju Wu (College of Life Sciences, Northeast Agricultural University, Harbin, China), was grown under natural sunlight on the campus of the Northeast Agricultural University (N 45°44'33.64", E 126°43'22.07") in Harbin, Heilongjiang Province of China. The samples, collected in September 2018, were dried, pulverized, and sifted through the 24 mesh sieve.

2.2. Preparation of *Syringa oblata* Lindl. for processing

The four kinds of processed *S. oblata* (leaves, 10 g each) were prepared in accordance with the Veterinary Pharmacopoeia of the People's Republic of China (volume II, 2015 edition) as follows.

In the process of stir-frying with vinegar, salt-water and ginger juice, the leaves of *S. oblata* were mixed well with 7 mL of

diluted rice vinegar, salt-water and ginger juice in a pot, respectively, and moistened for 15 min. They were then fried for 3 min and heated with a mild flame for 3 min at 400 °C until the vinegar, salt-water and ginger juice were completely absorbed. The ratios of 7 mL diluted rice vinegar, salt-water and ginger juice were 1 : 6, 0.2 : 7 and 1.25 : 7, respectively. In the process of stir-frying with honey, *S. oblata* was prepared by mixing it well with 7 mL of 50% refined honey in a pot, moistening for 3 h, frying for 3 min and heating with a mild fire for 3 min at 400 °C until the refined honey was completely absorbed. The refined honey was prepared by placing 20 g of honey in a pot, heating to slight boiling at 120 °C until the surface was light yellow and shiny uniform bubbles appeared, then filtering through cotton gauze. Finally, all of the processed *S. oblata* was removed and allowed to cool.

2.3. Determination of rutin and flavonoid content

A total of 20 to 50 mg of each type of processed *S. oblata* was used to extract flavonoids by adding 2 mL of 50% (v/v) methanol (HPLC grade) in H₂O. The mixture was placed in an ultrasonic cleaner for 20 min and centrifuged for 10 min at 13 000 rpm.²³ The supernatant was filtered through a 0.45 mm membrane filter and analyzed using a Waters Alliance HPLC system (Shimadzu, Corporation, Kyoto, Japan) and UV-1700 spectrophotometer (Shimadzu Corporation, Japan). Rutin and total flavonoids in normal and four kinds of processed *S. oblata* were determined using the chromatographic fingerprint analysis⁶ and aluminum chloride colorimetric method, respectively, as described previously.²⁴ Among them, the chromatogram and colorimetric values were monitored at a wavelength of 355 nm and 510 nm, respectively during the experiment.

2.4. Biofilm formation

In this study, the *S. xylosus* ATCC 700404 strain was cultured in trypticase soy broth (TSB; Summus Ltd., Harbin, Heilongjiang, China) at 37 °C for 12 h with constant shaking. The methanolic extracts of processed *S. oblata* were used for minimum inhibitory concentration (MIC) assays and crystal violet staining using the protocol described previously.²⁵ Rutin was used as a positive control.

2.5. Experimental design of one of the *Syringa oblata* Lindl. processes

In order to improve the rutin content in one of the best processed *S. oblata* samples, two parameters, including two kinds of vinegar (rice vinegar and vinegar) and moistening times (15 min, 30 min and 40 min), were screened using a single-factor experiment. On the basis of the single-factor experiment, the BBD is a useful statistical technique²⁶ that can be used to assess how the processing conditions influence the rutin and total flavonoid content. In this study, a three-level BBD experimental design required 17 experimental runs with three central points, which were employed to determine the optimal processing conditions. The selected independent variables were X_1 (moistening time), X_2 (frying time) and X_3 (the dosage of vinegar) (Table 1). Three levels were coded 1, 0 and -1



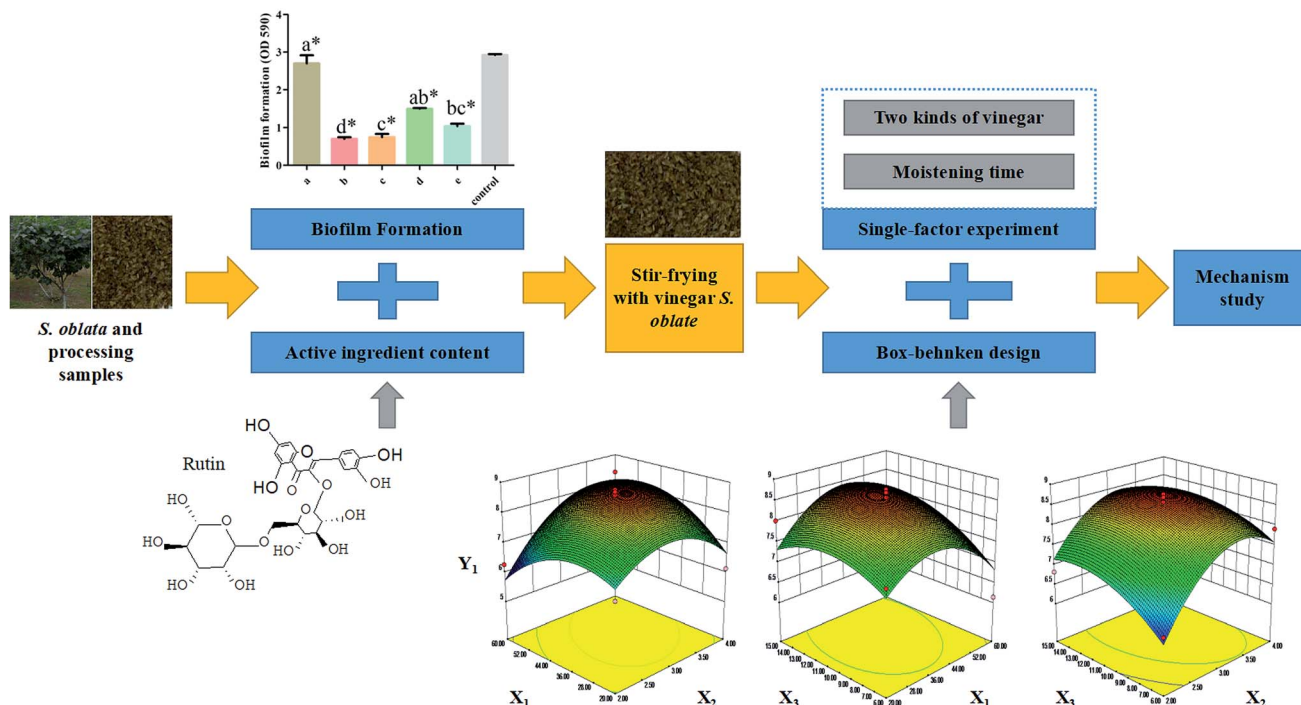


Fig. 1 The experimental flow chart.

for high, intermediate and low level, respectively (Table 1). The actual experimental design matrix is given in Table 2. Experimental data were analyzed by using a second-order polynomial model that correlated the interaction between the independent variables and the response variable.²⁷ The second-order polynomial equation is as follows:

$$Y = b_0 + \sum_{i=1}^k b_i X_i + \sum_{i=1}^k b_{ii} X_i^2 + \sum_{i < j} b_{ij} X_i X_j$$

where Y_1 ($\mu\text{g g}^{-1}$) and Y_2 (mg g^{-1}) are the predicted responses (the content of rutin and total flavonoids), X_i and X_j are the input variables (moistening time, frying time and the dosage of vinegar), b_0 , b_i , b_{ii} , and b_{ij} represent the regression coefficients of constant, linear, quadratic, and interactions terms, respectively, and k represents the number of variables.²⁸ According to the analysis of variance (ANOVA) with 95% confidence level, the regression coefficients, coefficient of determination, and the lack of fit for each studied response were employed to evaluate the fitness of the regression model. Design-expert software (version 8.0) was used to analyze the experimental data. The p -values < 0.05 were considered to be statistically significant.

2.6. Determination of the histidine content

Overnight cultures of *S. xyloso* were diluted with TSB (corresponding to 1×10^5 CFU mL^{-1}) and treated with sub-MICs processing *S. oblate*. This mixture was incubated at 37 °C for 24 h. Untreated *S. xyloso* served as the control. At the same time, the standard solution of 15 mg histidine (99% pure histidine was purchased from Beijing Solarbio Technology Co. Ltd.) was prepared by dissolving it in 250 mL (0.1 mol L^{-1})

hydrochloric acid. The above assay was conducted using HPLC external standard analysis.²⁵ The chromatographic separation was carried out on a Diamosil C18 column ($4.6 \text{ mm} \times 250 \text{ mm}$, $5 \mu\text{m}$) with a gradient solvent A (10 mmol L^{-1} diammonium hydrogen phosphate buffer containing 10 mmol L^{-1} sodium 1-octanesulfonate and obtaining pH 2.0 by phosphoric acid) and solvent B (acetonitrile) as the mobile phase at a flow rate of 1 mL min^{-1} . The gradient conditions were 0 to 5 min, 95% solvent A; 5 to 6 min, 95% to 86% solvent A; 6 to 15 min, 86% solvent A; 15 to 16 min, 86% to 87% solvent A; 16 to 25 min, 87% solvent A. The detection wavelength was set at 205 nm, and the injection volume was 100 μL . Three independent sample analyses were performed for each sample.

2.7. Verification of *hisB* expression in processing *Syring* *oblate* Lindl. by quantitative real-time PCR

In order to investigate the effect of sub-MICs of processed *S. oblate* on the expression of the IGP synthetic gene (*hisB*) of *S. xyloso*, the *S. xyloso* media culture (mid-log growth phase) was supplemented with sub-MICs of processed *S. oblate* and

Table 1 Variables and experimental design levels for response surface

Independent variable/unit	Coded symbol	Level		
		−1	0	1
Moistening time/min	X_1	20	40	60
Frying time/min	X_2	2	3	4
The dosage of vinegar/mL	X_3	6	10.5	15

Table 2 Box–Behnken design matrix with three independent variables expressed in code with experimental and predicted responses

Run	X_1	X_2	X_3	Y_1 ($\mu\text{g g}^{-1}$)		Y_2 (mg g^{-1})	
				Predicted value (%)	Experiment value (%)	Predicted value (%)	Experiment value (%)
1	1	1	0	8.37	8.00	306.06	287.84
2	2	0	−1	6.13	6.83	249.31	260.20
3	0	0	0	8.57	8.55	311.98	310.02
4	−1	−1	0	6.56	6.92	242.52	260.74
5	0	1	1	7.64	7.80	278.47	281.01
6	0	0	0	8.76	8.55	311.98	310.02
7	−1	0	1	8.01	7.32	288.35	277.46
8	0	1	−1	7.90	7.58	263.97	271.29
9	3	0	1	7.71	7.91	257.01	272.69
10	−1	1	0	6.13	6.66	270.16	278.51
11	0	0	0	8.68	8.55	311.98	310.02
12	0	−1	1	6.82	7.15	270.75	263.43
13	4	−1	0	6.27	5.74	261.74	253.39
14	0	0	0	8.42	8.55	311.98	310.02
15	0	−1	−1	6.39	6.23	239.18	236.65
16	−1	0	−1	7.46	7.26	269.14	253.46
17	0	0	0	8.34	8.55	302.15	310.02

incubated with shaking at 37 °C for 24 h. Cells without sub-MICs of processed *S. oblata* served as the control. The supplemented solution was centrifuged at $10\,000 \times g$ for 5 min and treated with an RNASE REMOVER I (Huayueyang Ltd., Beijing, China). The E.Z.N.A.TM bacterial RNA isolation kit was used to determine the total RNA levels. The 16S RNA gene was selected as an internal control. The specific primers were obtained from Sangon Biotech (Shanghai) and listed in Table S1.† Quantitative real-time PCR was performed as described in our previous study.²⁹

2.8. Statistical analysis

All the experiments were performed in triplicate. Data analysis and calculation of the standard deviation were done using SPSS 11.0.0 (IBM, USA) and $p < 0.05$ was considered indicative of statistically significant differences.

3. Results

3.1. Rutin detection in processed *Syringa oblata* Lindl.

To investigate the effect of rutin in the normal and the four processed *S. oblata* samples, an investigation was conducted as described in the Materials and Methods section above. Previous studies have shown that at any given temperature, the longer the herb was subjected to stir-frying, the less rutin was extracted.³⁰ When the heating temperature was 180 °C to 200 °C, rutin was partially decomposed to quercetin and was entirely damaged at temperatures above 210 °C. This is also in tandem with our study that when the heated temperature exceeded 200 °C, the rutin content in all the processed *S. oblata* was significantly lower ($p < 0.05$) than the normal *S. oblata* (Fig. 2A). At the same time, it may be beneficial to extract rutin with acid at a suitable pH (pH 2.0), which is consistent with the pH of rice vinegar.³¹ Thus, this study showed that the content of rutin in

the stir-frying with vinegar was the highest among all the processed *S. oblata*. Chromatograms of rutin and *S. oblata* in the different processing groups are shown in Fig. S1.†

3.2. Flavonoid detection in processed *Syringa oblata* Lindl.

To investigate the effect of rutin in normal and the four processed *S. oblata* samples, an investigation was done as described in the Materials and Methods section above. This study showed that the total flavonoid contents of the *S. oblata* samples stir-fried with salt-water, ginger juice and honey were 115.92 ± 5.88 , 115.59 ± 4.84 and 93.41 ± 3.91 mg g^{-1} , respectively. This indicated that the concentration of total flavonoids in *S. oblata* improved and was slightly lower than the normal *S. oblata* (116.86 ± 1.14 mg g^{-1}) ($p < 0.05$) (Fig. 2B). This is in consonance with the previous study wherein the total flavonoid content decreased after stir-frying.³² In addition to the above-mentioned processing method, it has been reported that the total flavonoids content in *S. oblata* processed with vinegar increased by 21% in comparison with unprocessed *S. oblata*.³³ This study has proven that the stir-frying *S. oblata* with vinegar had the strongest effect and the concentration of total flavonoids was 121.67 ± 2.08 mg g^{-1} and significantly higher than the normal *S. oblata* ($p < 0.05$).

3.3. Biofilm inhibition properties of processed *Syringa oblata* Lindl.

The normal and the four processed *S. oblata* samples were evaluated against *S. xylosus* biofilm formation with the same MICs (31.25 mg mL^{-1}). The results revealed that compared with the control, 1/2 MIC of the normal and four processed *S. oblata* had the significant ability to inhibit *S. xylosus* biofilm formation ($p < 0.05$) (Fig. 2C). Most recent studies have shown that *S. oblata* can inhibit *Streptococcus suis* (*S. suis*) biofilm formation



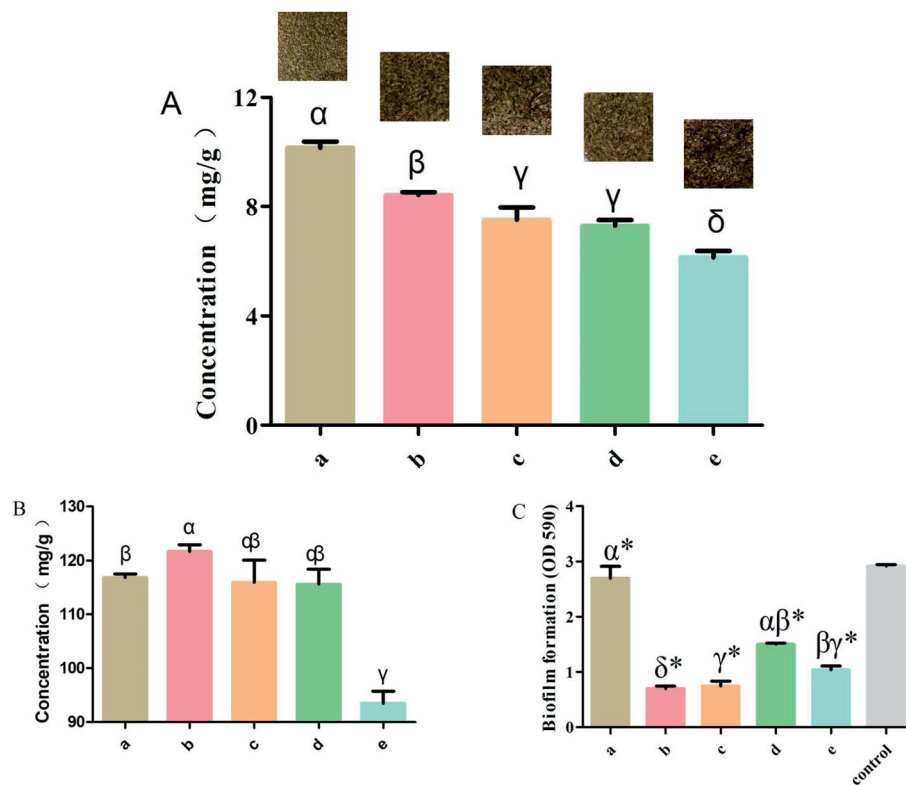


Fig. 2 (A) The rutin content with different *Syringa oblata* Lindl. processing methods. (B) The total flavonoid content with different *Syringa oblata* Lindl. processing methods. (C) The inhibition of *Staphylococcus xylosus* biofilm formation by different *Syringa oblata* Lindl. processing methods: (a) normal, (b) stir-frying with vinegar, (c) stir-frying with salt-water, (d) stir-frying with ginger juice and (e) stir-frying with honey. Data are expressed as mean \pm standard deviation ($n = 3$). Different letters indicate a significant difference at $p < 0.05$. (* $p < 0.05$) is significantly different as compared to the untreated control bacteria.

and rutin is the active ingredient.⁶ Interestingly, we have shown here that the *S. oblata* stir-fried with vinegar had the greatest ability to inhibit the formation of *S. xylosus* biofilm (Fig. 2C). At the same time, the rutin and total flavonoid contents in the *S. oblata* stir-fried with vinegar were the highest among all processed *S. oblata*. Thus, in order to improve the ability of *S. xylosus* to form biofilm, the rutin and flavonoid content needed to be improved by optimizing the *S. oblata* stir-fried with vinegar.

3.4. Box-Behnken design and response surface methodology

3.4.1. Single-factor experiments for the *Syringa oblata* Lindl. stir-fried with vinegar. The types of vinegar (rice vinegar and vinegar), the most widely used adjuvants,⁴ and moistening time (15 min, 30 min and 40 min) are among the important factors affecting the quality of the *S. oblata* stir-fried with vinegar.³⁴ The results showed that on adding the 7 mL diluted rice vinegar (1 : 6 ratio), moistening for 15 min and frying for 2 min, the rutin of the *S. oblata* stir-fried with rice vinegar was significantly higher than the *S. oblata* stir-fried with vinegar ($p < 0.05$) (Fig. 3A). As the moistening time increased from 15 min to 40 min, the rutin of the *S. oblata* stir-fried with rice vinegar increased first and then decreased (Fig. 3B). Thus, the rice vinegar and 30 min moistening time were chosen as the central point of the BBD experiment.

3.4.2. Optimization of the *Syringa oblata* Lindl. stir-fried with vinegar by the Box-Behnken design. To evaluate all the identified parameters together, the BBD experiment with a total of 17 runs, based on the effects of the most impacting factors, was used to optimize the processing conditions of the *S. oblata* stir-fried with rice vinegar based on the single-factor experiments. The design matrix and corresponding results of the RSM experiments to evaluate the three independent variables, including X_1 , X_2 and X_3 with experimental and predicted responses, are shown in Table 2. The response values for Y_1 and Y_2 ranged from 6.13 to 8.48 $\mu\text{g g}^{-1}$ and 239.18 to 311.98 mg g^{-1} , respectively.

By applying multiple regression analysis to the experimental data, the relationship between the response variable and the test variables was described by the following second-order polynomial equation:

$$Y_1 = 8.55 + 0.5X_1 + 0.63X_1X_2 - 0.93X_1^2 - 0.79X_2^2 \quad (1)$$

$$Y_2 = 310.02 + 13.06X_1 - 21.38X_1^2 - 18.52X_2^2 - 25.54X_3^2 \quad (2)$$

The results from the ANOVA for the adjusted model for Y_1 and Y_2 are presented in Table 3. The model F -values of 4.28 and 4.06, both with p -value < 0.05 , and the lack of fit F -values of 23.22 and 28.48, imply that the model was significant and well



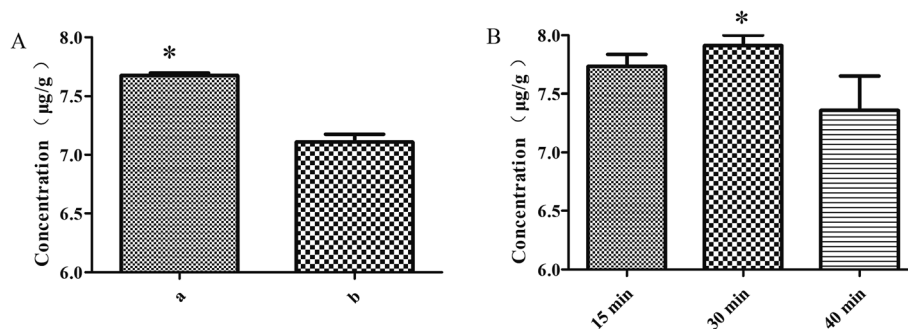


Fig. 3 Single-factor experiments with the rice vinegar (a) and vinegar (b) (A), and moistening time (B) in the *Syringa oblata* Lindl. stir-fried with vinegar. Data are expressed as mean \pm standard deviation ($n = 3$).

adapted to the response. The results showed that the determination coefficient (R^2) was 0.8464 and 0.8392, which were in good agreement with the adjusted R_{adj}^2 of 0.6488 and 0.6323. It is also indicated that the model was highly significant. A relatively lower value of the coefficient of variation (CV = 7.54% and 5.63%) indicated the better precision and reliability of the experiments carried out and the selected quadratic models exhibited high accuracy and significant reproducibility.³⁵

Response surface plots for each response condition and composite desirability functions were also analyzed to optimize the formulation. The optimized conditions were 44.03 and 40.68 moisturizing time (min), 3.3 and 3.29 frying time (min) and 11.97 and 11.19 vinegar dosage (mL) with $8.68 \mu\text{g mg}^{-1}$ and 312.64 mg g^{-1} calculated as the predicted values of Y_1 (Fig. 4) and Y_2 (Fig. S2†), respectively. This similarity indicated that the processing regression model was highly significant and could be used to predict the concentration of *S. oblata*.

3.5. Verification of the predictive model by biofilm formation

The optimum conditions obtained by RSM were used to validate the predictive model for the processing conditions of the *S.*

oblata stir-fried with rice vinegar as compared with normal *S. oblata* by inhibiting biofilm formation. The results revealed that compared with the control, 1/2 MIC of the *S. oblata* stir-fried with vinegar had the highest ability to inhibit the formation of *S. xyloso* biofilm as compared to 1/2 MIC of normal *S. oblata* and a positive control rutin (Fig. 5A). Therefore, the response model was adequate for reflecting the expected optimization, and the extraction conditions and the ability to inhibit biofilm formation achieved by RSM were reliable and practical.

3.6. Determination of histidine content

The histidine content of *S. xyloso* was determined in sub-MICs processed *S. oblata* in the normal and the stir-fried with vinegar group. When supplemented with sub-MICs of processed *S. oblata* in the culture medium, the histidine content of *S. xyloso* showed significant depletion ($p < 0.05$) in comparison with the control (Fig. 5B). Additionally, the histidine content of *S. xyloso* in the *S. oblata* stir-fried with vinegar *S. oblata* was significantly lower than in normal *S. oblata* ($p < 0.05$) (Fig. 5B). At the same time, the chromatograms of histidine, the control group without *S. oblata*, normal *S. oblata* and the *S. oblata* stir-fried with vinegar are shown in Fig. S3.†

Table 3 ANOVA for the response surface quadratic model of the *Syringa oblata* Lindl. stir-fried with vinegar^a

Source	Sum of squares		df		Mean square		F value		p-value	
	a	b	a	b	a	b	a	b	a	b
Model	12.45	9024.24	9	9	1.38	1002.69	4.28	4.06	0.034*	0.039*
X_1	2.01	1363.95	1	1	2.01	1363.95	6.22	5.52	0.041*	0.051
X_2	0.01	1.96	1	1	0.01	1.96	0.04	7.92×10^{-3}	0.851	0.932
X_3	0.66	665.85	1	1	0.66	665.85	2.03	2.69	0.197	0.145
X_1X_2	1.59	69.60	1	1	1.59	69.60	4.94	0.28	0.062	0.612
X_1X_3	0.12	72.77	1	1	0.12	72.77	0.37	0.29	0.561	0.604
X_2X_3	0.26	33.12	1	1	0.26	33.12	0.82	0.13	0.396	0.725
X_1^2	3.64	1924.09	1	1	3.64	1924.09	11.28	7.79	0.012*	0.027*
X_2^2	2.63	1443.72	1	1	2.63	1443.72	8.13	5.84	0.025*	0.046*
X_3^2	0.79	2747.37	1	1	0.79	2747.37	2.45	11.12	0.162	0.013*
Residual	2.26	1729.78	7	7	0.32	247.11				
Lack of fit	2.14	1652.42	3	3	0.71	550.81	23.22	28.48	0.005*	0.004*
Pure error	0.12	77.36	4	4	0.03	19.34				
Cor total	14.71	10 754.02	16	16						

^a *Significant ($p < 0.05$); a indicated rutin content. b indicated total flavonoid content.



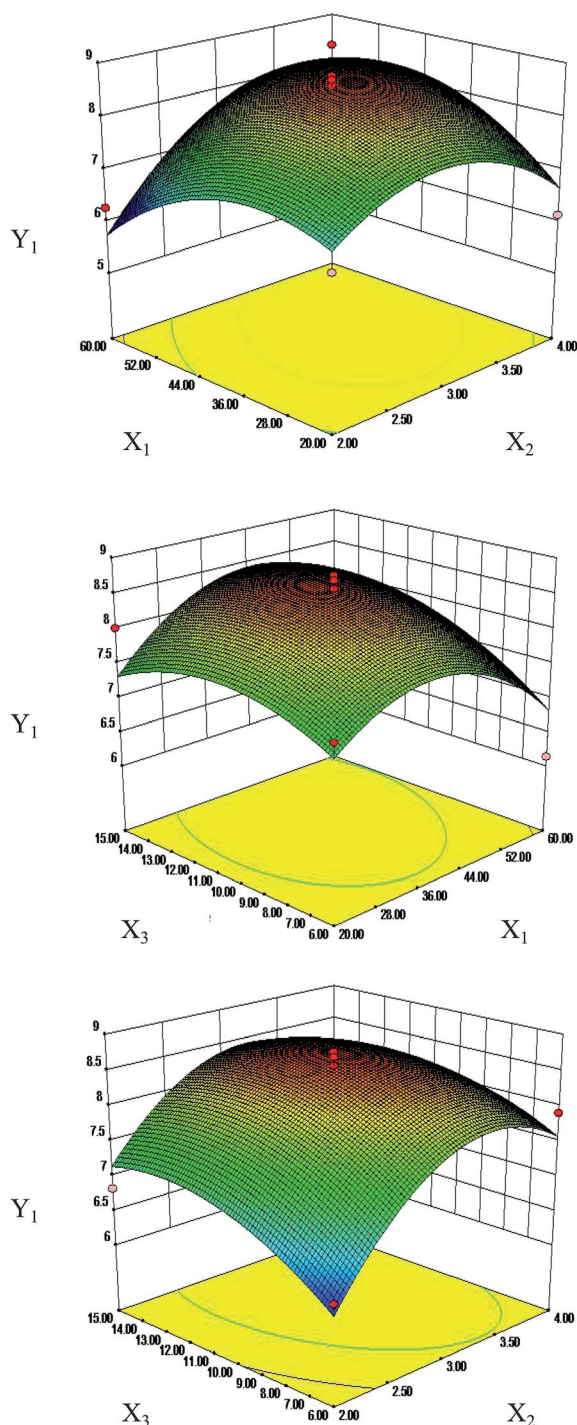


Fig. 4 Response surface plots (three-dimensional) showing the effects of the moistening time (X_1 , min), frying time (X_2 , min) and the dosage of vinegar (X_3 , mL) on the concentration of rutin (Y_1 , $\mu\text{g g}^{-1}$) and total flavonoids (Y_2 , mg g^{-1}) in *Syringa oblata* Lindl.

3.7. Verification of *hisB* expression in processed *Syringa oblata* Lindl. by quantitative real-time PCR

To further investigate the expression of *hisB* during the interference of the normal *S. oblata* and the *S. oblata* stir-fried with

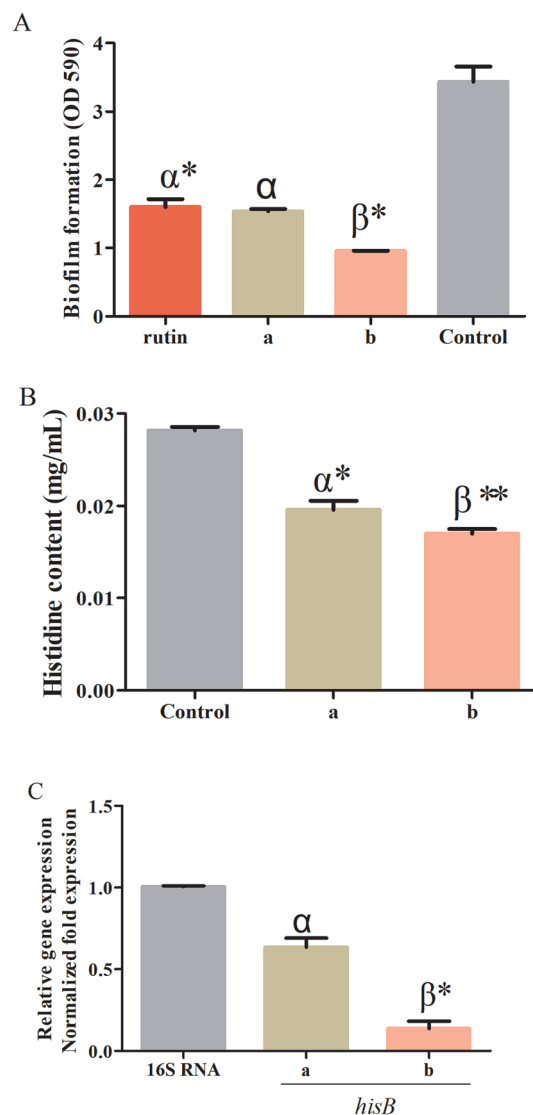


Fig. 5 (A) The inhibition of *Staphylococcus xylosus* biofilm formation by rutin: normal (a) and stir-frying with vinegar (b). (B) Determination of the histidine content and of *Staphylococcus xylosus* in sub-MICs in the control, normal *Syringa oblata* Lindl. (a) and *Syringa oblata* Lindl. stir-fried with vinegar (b). (C) The effect of *Syringa oblata* Lindl. on the expression of *hisB* genes in *Staphylococcus xylosus*. The expression was normalized to that of 16S RNA. Controls refer to the absence of *Syringa oblata* Lindl. * $p < 0.05$ and ** $p < 0.01$ as compared to the control bacteria. Different letters indicate a significant difference at $p < 0.05$. Data are expressed as mean \pm standard deviation ($n = 3$).

vinegar with the formations of *S. xylosus* biofilms, the sub-MICs were determined by quantitative real-time PCR. The results showed that the sub-MICs of the normal and the *S. oblata* stir-fried with vinegar interfered with the formation of biofilms by *S. xylosus*, and the expression of *hisB* gene was significantly down-regulated ($p < 0.05$ and $p < 0.01$) (Fig. 5C). Compared with the sub-MICs of the normal *S. oblata*, the expression of *hisB* gene was significantly down-regulated with the sub-MICs of the *S. oblata* stir-fried with vinegar.



4. Discussion

Herbal medicine processing is a traditional herbal pharmaceutical technology.³⁶ There are complex chemical changes during herbal medicine processing, and these chemical constituents that change during processing may be the basis for relieving drug irritation, enhancing therapeutic effect, biological efficacy, clinical applicability³⁷ and clinical efficacy changes.³⁶

In recent years, there have been many studies on the chemical and molecular mechanisms involved in the processing of herbal medicine.³⁶ For example, in stir-fried-treated *Adiandra nitida* tea, a well-known resource of functional foods, the main compounds such as total water-soluble solids, phenolics and flavonoids content were higher than in the corresponding untreated samples.³⁸ Bai-Zhu possesses significantly different therapeutic effects as compared to processed Bai-Zhu since the processing increases the content of atractylenolide, which could strengthen the effect on gastrointestinal function.³ There are no studies on the processing of *S. oblata* at home and abroad. This study has demonstrated that stir-frying *S. oblata* with vinegar led to the highest content of rutin and total flavonoids among all processed *S. oblata*.

At present, the RSM experimental designs most commonly applied in pharmaceutical experiments are the central composite design (CCD) and Box-Behnken design (BBD).³⁹ Compared to CCD, BBD employs a reduced number of experimental runs, leading to higher efficiency³⁹ that is obtained with the best conditions of resources²¹ for quadratic models,²¹ and it is extensively used in RSM for three-level factors.²⁰ BBD has been used to optimize the extraction of bioactive and antibacterial compounds from different processing samples. Thus, it is necessary to improve the content of rutin and biofilm formation in the *S. oblata* stir-fried with vinegar by the RSM method. This study has confirmed that the optimized conditions were rice vinegar, 44.03 min moisturizing time, 3.3 min frying time and 11.97 mL vinegar dosage, with rutin content calculated as 8.68 $\mu\text{g mg}^{-1}$ for the single-factor and BBD experiments with *S. oblata* stir-fried with vinegar. When the main factors were the dosage of rice vinegar, vinegar infiltration, and processing temperature, with values of 20%, 3 h and 130, the contents of effective components were increased.⁴⁰

We have demonstrated that the *S. oblata* stir-fried with vinegar had the greatest ability to inhibit the formation of *S. xylosus* biofilm as compared to the normal *S. oblata*. A study has found that the nitrogen metabolic limitation has a key role in the formation of biofilm, namely, histidine metabolism.¹³ Therefore, it was necessary to investigate the mechanisms underlying the effects of *S. oblata* stir-fried with vinegar on histidine metabolism and biofilm formation. The results showed that compared with normal *S. oblata*, the histidine content and *hisB* expression of *S. xylosus* in the *S. oblata* stir-fried with vinegar were significantly decreased and down-regulated. This is consistent with previous research where L-histidine dramatically decreased the biofilm formation by *Saccharomyces cerevisiae* flor yeasts.¹⁴ This may prove that the *S.*

oblata stir-fried with vinegar may indirectly interfere with *S. xylosus* biofilm formation by decreasing histidine content and *hisB* gene expression.

5. Conclusion

In conclusion, we successfully developed a method for processing *S. oblata*, namely, stir-frying the *S. oblata* with vinegar, which was sufficient to inhibit biofilm formation. The mechanism related to the biofilm formation of *S. xylosus* involved the decrease and down-regulation of the histidine content and *hisB* gene expression, respectively, upon treatment with *S. oblata* stir-fried with vinegar as compared with normal *S. oblata*. These results provide an interesting template for designing new and more effective processing methods.

Author contributions

Y-HL designed the whole experiment; Y-YL and X-RC directed the completion of the experiment; Y-HZ, X-XX, Q-WQ, X-YC, W-YD, G-LC, A-JW, X-WF, GB, EN were supportive during the experiment.

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Conflicts of interest

The authors declare no competing financial interest.

References

- 1 T. Yi, J.-Y. Fang, L. Zhu, Y.-N. Tang, H. Ji, Y.-Z. Zhang, J.-C. Yu, X.-J. Zhang, Z.-L. Yu, Z.-Z. Zhao and H.-B. Chen, *Chin. Med.*, 2016, **11**, 1–8.
- 2 W. Wang, J. Xue, K. Li, D. Hu, G. Huang and L. Ye, *Int. J. Food Prop.*, 2017, **20**, S644–S653.
- 3 G.-S. Shan, L.-X. Zhang, Q.-M. Zhao, H.-B. Xiao, R.-J. Zhuo, G. Xu, H. Jiang, X.-M. You and T.-Z. Jia, *J. Pharm. Biomed. Anal.*, 2014, **98**, 74–84.
- 4 J.-Y. Fang, L. Zhu, T. Yi, J.-Y. Zhang, L. Yi, Z.-T. Liang, L. Xia, J.-F. Feng, J. Xu, Y.-N. Tang, Z.-Z. Zhao and H.-B. Chen, *Chin. Med.*, 2015, **10**, 1–7.
- 5 S. Guozhu, C. Yuan, L. Chun, Y. Xuelong, G. Xiaoli, T. Pengfei and C. Xingyun, *Chem. Cent. J.*, 2015, **9**, 1–12.
- 6 Y.-Y. Liu, X.-R. Chen, L.-F. Gao, M. Chen, W.-Q. Cui, W.-Y. Ding, X.-Y. Chen, B.-O. God'spower and Y.-H. Li, *Front. Pharmacol.*, 2018, **9**, 1–10.
- 7 J. Bai, Y. Yang, S. Wang, L. Gao, J. Chen, Y. Ren, W. Ding, I. Muhammad and Y. Li, *Front. Pharmacol.*, 2017, **8**, 1–9.
- 8 K. M. Osman, K. A. Abd El-Razik, H. S. H. Marie and A. Arafa, *J. Food Saf.*, 2016, **36**, 89–99.



- 9 B. Kot, M. Piechota, M. Antos-Bielska, E. Zdunek, K. M. Wolska, T. Binek, J. Olszewska, P. Gulinski and E. A. Trafny, *Pol. J. Vet. Sci.*, 2012, **15**, 741–749.
- 10 S. Planchon, M. Desvaux, I. Chafsey, C. Chambon, S. Leroy, M. Hébraud and R. Talon, *J. Proteome Res.*, 2009, **8**, 1797.
- 11 S. Planchon, B. Gaillardmartinie, E. Dordetfrisoni, M. N. Bellonfontaine, S. Leroy, J. Labadie, M. Hébraud and R. Talon, *Int. J. Food Microbiol.*, 2006, **109**, 88–96.
- 12 R. K. Kulis-Horn, M. Persicke and J. Kalinowski, *Microb. Biotechnol.*, 2014, **7**, 5–25.
- 13 C. G. Xu, Y. B. Yang, Y. H. Zhou, M. Q. Hao, Y. Z. Ren, X. T. Wang, J. Q. Chen, I. Muhammad, S. Wang and D. Liu, *Front. Pharmacol.*, 2017, **8**, 543.
- 14 M. B. Zeidan, G. Zara, C. Viti, F. Decorosi, I. Mannazzu and M. Budroni, *PLoS One*, 2015, **10**, 1–10.
- 15 M. A. Bezerra, R. E. Santelli, E. P. Oliveira, L. S. Villar and L. A. Escaleira, *Talanta*, 2008, **76**, 965–977.
- 16 C. Liyana-Pathirana and F. Shahidi, *Food Chem.*, 2005, **93**, 47–56.
- 17 J. Carlos Martinez-Patino, B. Gullon, I. Romero, E. Ruiz, M. Brncic, J. S. Zlabur and E. Castro, *Ultrason. Sonochem.*, 2019, **51**, 487–495.
- 18 S. V. F. Gomes, L. A. Portugal, J. P. dos Anjos, O. N. de Jesus, E. J. de Oliveira, J. P. David and J. M. David, *Microchem. J.*, 2017, **132**, 28–35.
- 19 P. Naderi, M. Shirani, A. Semnani and A. Goli, *Ecotoxicol. Environ. Saf.*, 2018, **163**, 372–381.
- 20 V. Londhe and R. Shirsat, *AAPS PharmSciTech*, 2018, **19**, 1392–1400.
- 21 F. G. B. Los, A. A. F. Zielinski, J. P. Wojeicchowski, A. Nogueira and I. M. Demiate, *Food Anal. Methods*, 2019, **12**, 148–159.
- 22 B. Singh, H. K. Sharma and B. C. Sarkar, *J. Food Sci. Technol.*, 2012, **49**, 294–308.
- 23 S. Czermmel, J. Hoell, R. Loyola, P. Arce-Johnson, J. Antonio Alcalde, J. Tomas Matus and J. Bogs, *Front. Plant Sci.*, 2017, **8**, 1–15.
- 24 Z.-L. Sheng, P.-F. Wan, C.-L. Dong and Y.-H. Li, *Ind. Crops Prod.*, 2013, **43**, 778–786.
- 25 Y.-h. Zhou, C.-g. Xu, Y.-b. Yang, X.-x. Xing, X. Liu, Q.-w. Qu, W.-y. Ding, G. s. Bello-Onaghise and Y.-h. Li, *Front. Microbiol.*, 2018, **9**, 1–10.
- 26 S. Teixeira, C. Delerue-Matos and L. Santos, *Sci. Total Environ.*, 2019, **646**, 168–176.
- 27 P. Sharma, L. Singh and N. Dilbaghi, *J. Hazard. Mater.*, 2009, **164**, 1024–1029.
- 28 D. C. Montgomery, *Environ. Prog. Sustainable Energy*, 2013, **32**, 8–10.
- 29 W. Ding, Y. Zhou, Q. Qu, W. Cui, B. O. God'spower, Y. Liu, X. Chen, M. Chen, Y. Yang and Y. Li, *Front. Pharmacol.*, 2018, **9**, 1–11.
- 30 I. M. Liu and S. J. Sheu, *Am. J. Chin. Med.*, 1989, **17**, 179–187.
- 31 A. Tang, L. Ning and N. Cheng, *Acta Med. Sin.*, 1994, 1–4.
- 32 Q. Zhang, Y.-L. Wang, D. Gao, L. Cai, Y.-Y. Yang, Y.-J. Hu, F.-Q. Yang, H. Chen and Z.-N. Xia, *Pharm. Biol.*, 2018, **56**, 67–75.
- 33 K. Tang, W. Q. Wang and L. I. Biao, *Chin. J. Exp. Tradit. Med. Formulae*, 2013, 1–3.
- 34 N. U. o. T. C. M. N. L. Tulin, C. Mao and D. Ye, Optimization of Processing Technique of Sparganium stoloniferum by Orthogonal Experimental Design, *Chin. Tradit. Herb. Drugs*, 2000, 1–3.
- 35 G. D. Vyavahare, R. G. Gurav, P. P. Jadhav, R. R. Patil, C. B. Aware and J. P. Jadhav, *Chemosphere*, 2018, **194**, 306–315.
- 36 Y. Q. Xiao, L. Li, Y. Liu, Y. L. Ma and D. R. Yu, *China J. Chin. Mater. Med.*, 2016, **41**, 24–27.
- 37 Y. Y. Xu, H. Cai, G. Cao, Y. Duan, K. Pei, S. C. Tu, J. Zhou, L. Xie, D. D. Sun, J. Y. Zhao, J. Liu, X. Q. Wang and L. Shen, *J. Chromatogr., B*, 2018, **1083**, 110–123.
- 38 Y. S. Chen, Y. B. Shen, X. Fu, A. M. Abbasi and R. Yan, *Int. J. Food Sci. Technol.*, 2017, **52**, 1820–1827.
- 39 J. W. Tak, B. Gupta, R. K. Thapa, K. B. Woo, S. Y. Kim, T. G. Go, Y. Choi, J. Y. Choi, J.-H. Jeong, H.-G. Choi, C. S. Yong and J. O. Kim, *AAPS PharmSciTech*, 2017, **18**, 1125–1134.
- 40 J. Lei, Y. Cui and L. Fang, *Shaanxi J. Tradit. Chin. Med.*, 2015, 1–5.

